

Power law estimation

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Why look for power laws ?

■ Basic property : scale invariance $y(\lambda t) \sim \lambda^k y(t)$

In the spectral domain $\tilde{y}(f) \propto f^{-\beta}$

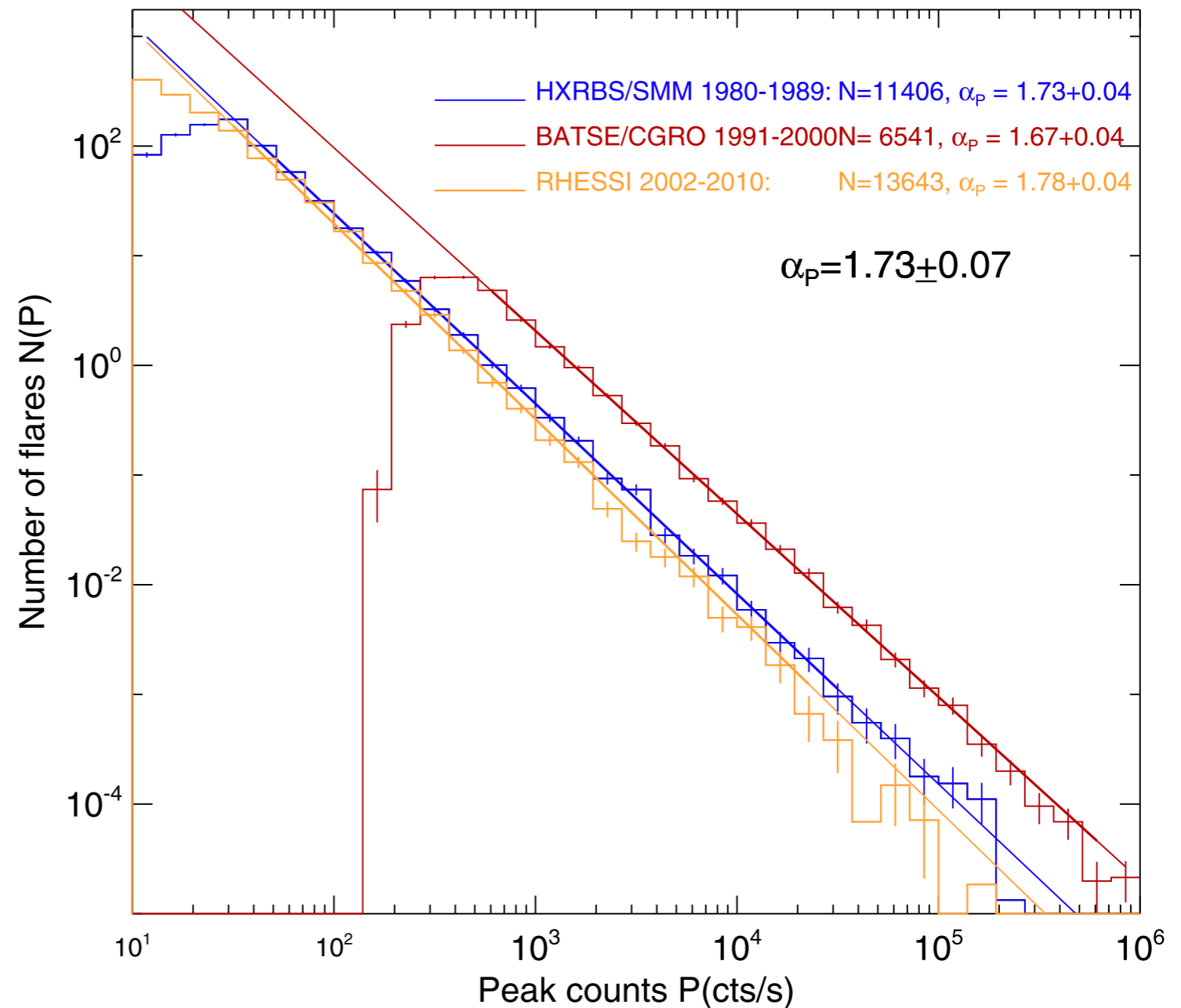


Why look for power laws ?

- Lots of fascinating properties
 - scale-invariance (no characteristic scale)
 - large deviations (rare events)
 - turbulence (intermittency, energy cascades, ...)
- Self-similarity is ubiquitous in plasmas
 - solar flare energy distribution
 - plasma turbulence
 - bursty bulk flows in the magnetosphere
 - photospheric magnetic field
 - ...

Example: flare energy distribution

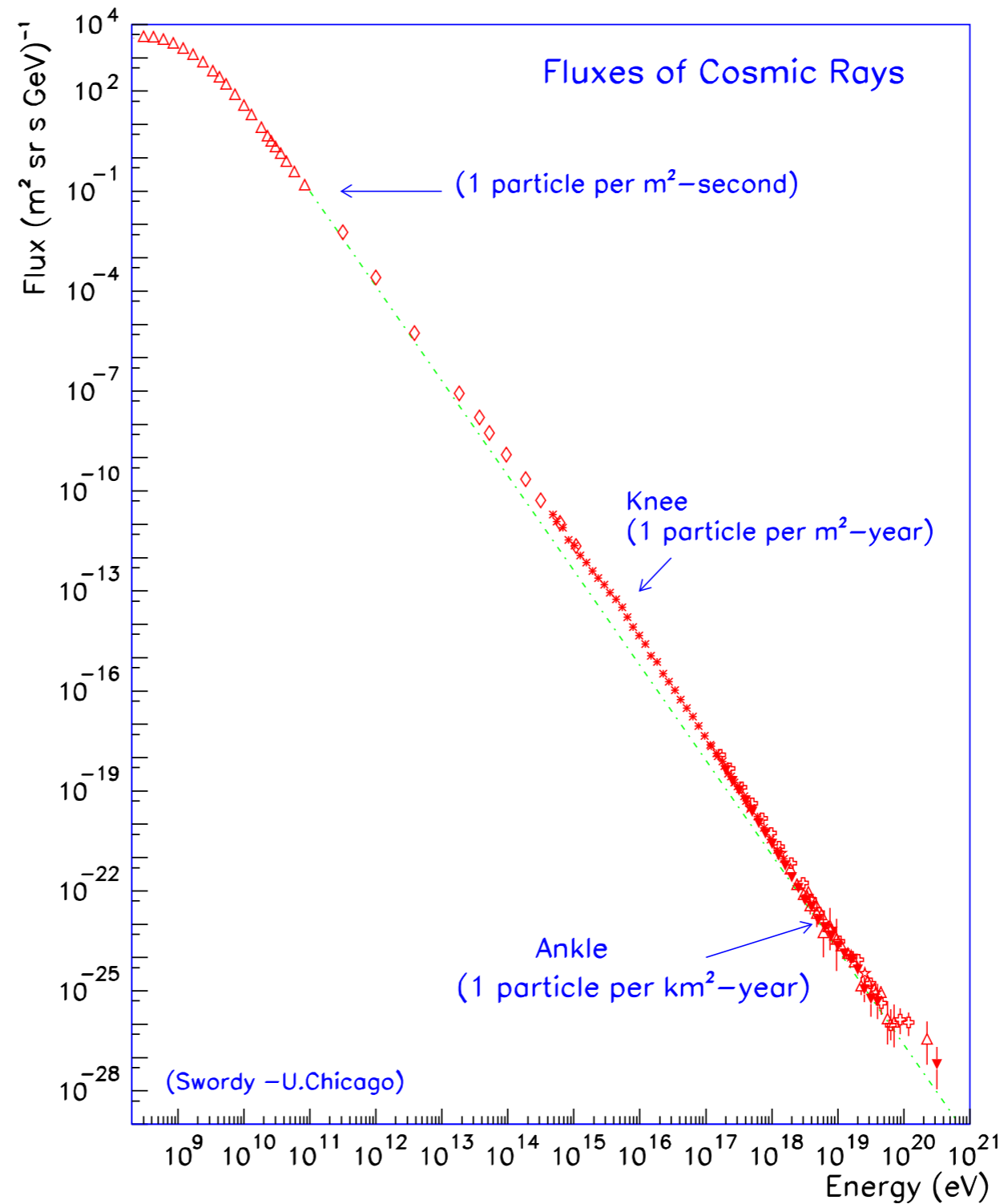
Occurrence frequency distributions of hard X-ray peak count rates from 3 instruments.



Aschwanden et al., 2014

Example: cosmic ray spectrum

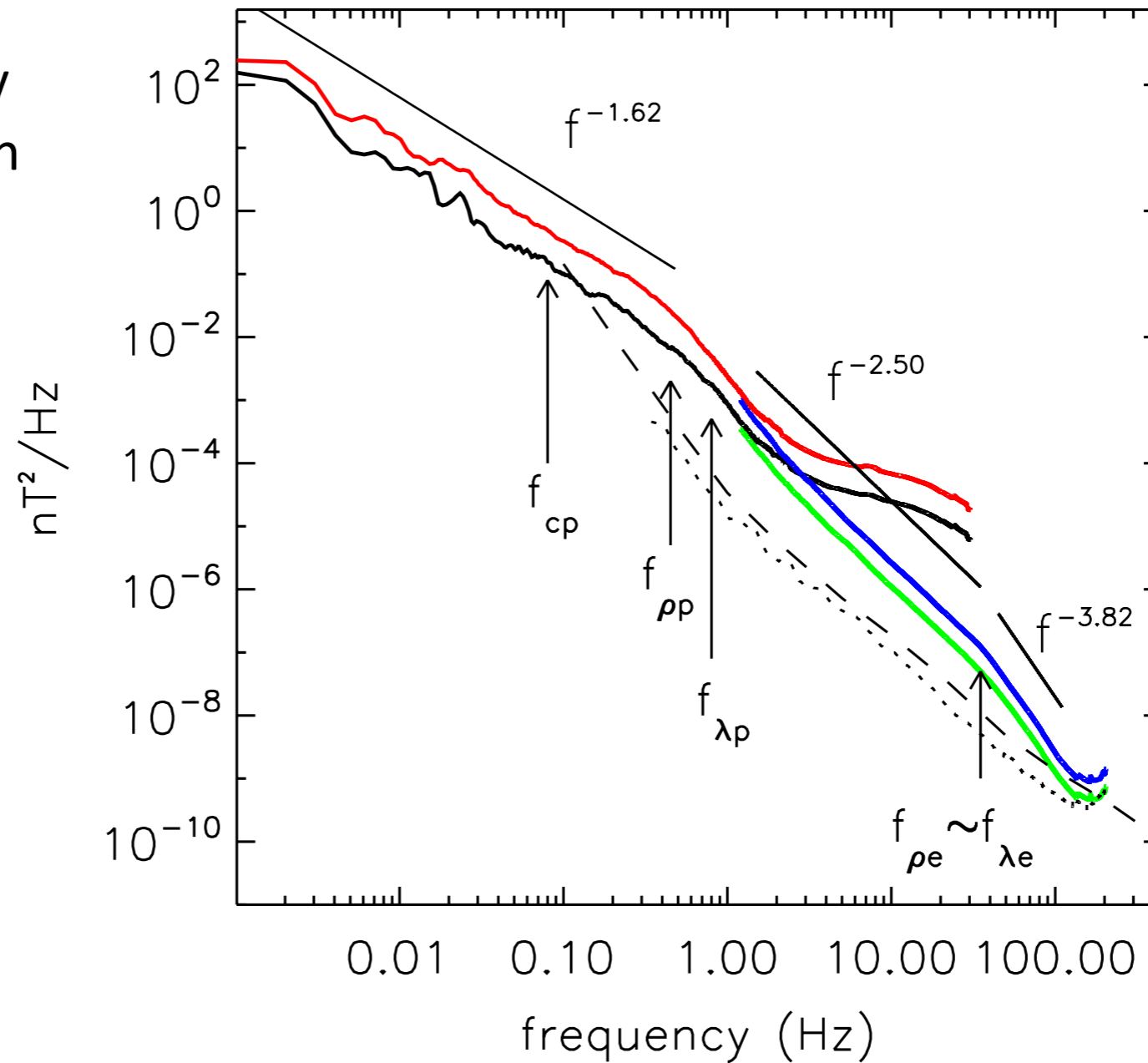
Cosmic ray spectrum between 1 GeV and 1 ZeV



Aschwanden et al., 2014

Example: plasma turbulence

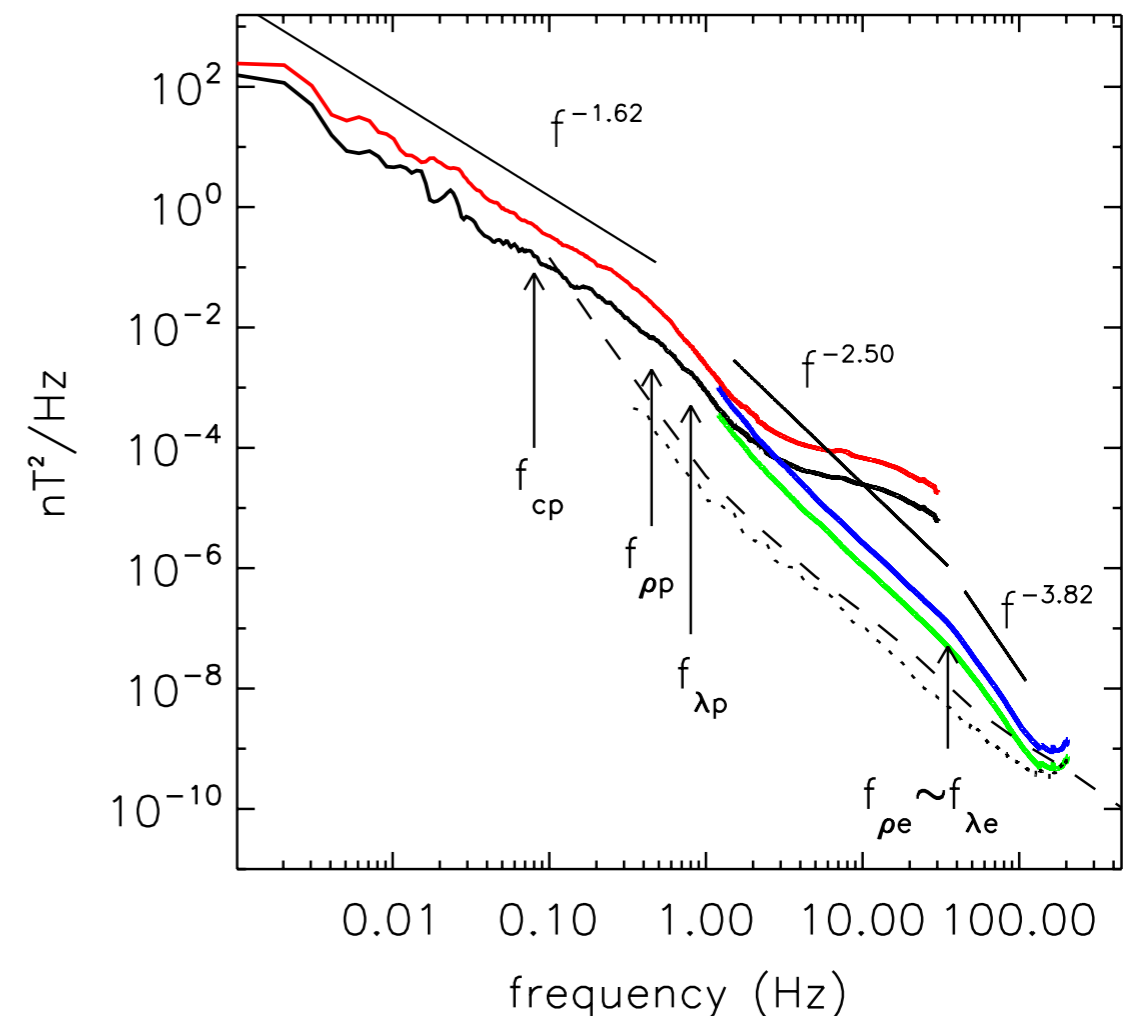
Power spectral density
of the magnetic field in
solar wind turbulence



Sahraoui et al., 2009

Why look for power laws ?

- The power law itself is **totally uninteresting**
- What physical process gave rise to it ?
 - where does the scaling start ?
 - where does the scaling end ?
 - what is its slope ?





Example :
solar wind magnetic field
recorded by Anik F1R

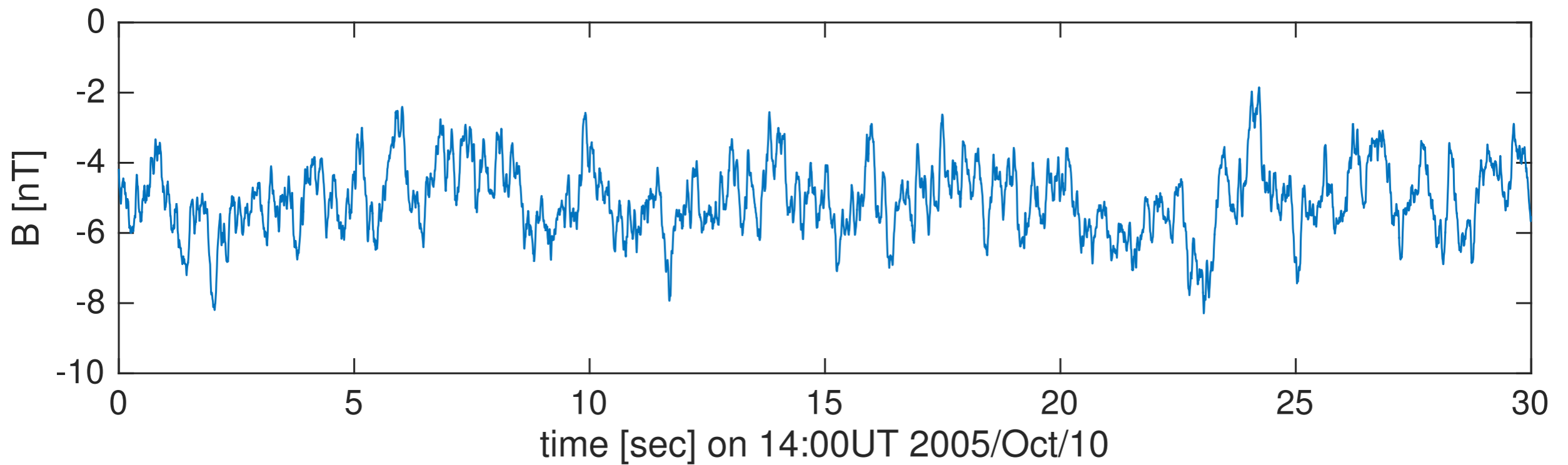
Solar wind observations

- Anik F1R (2005-today): Canadian geostationary satellite
 - experimental magnetic field sensor (fluxgate, 1 component)
 - high cadence : < 200 samples / second
 - some observations made in the solar wind



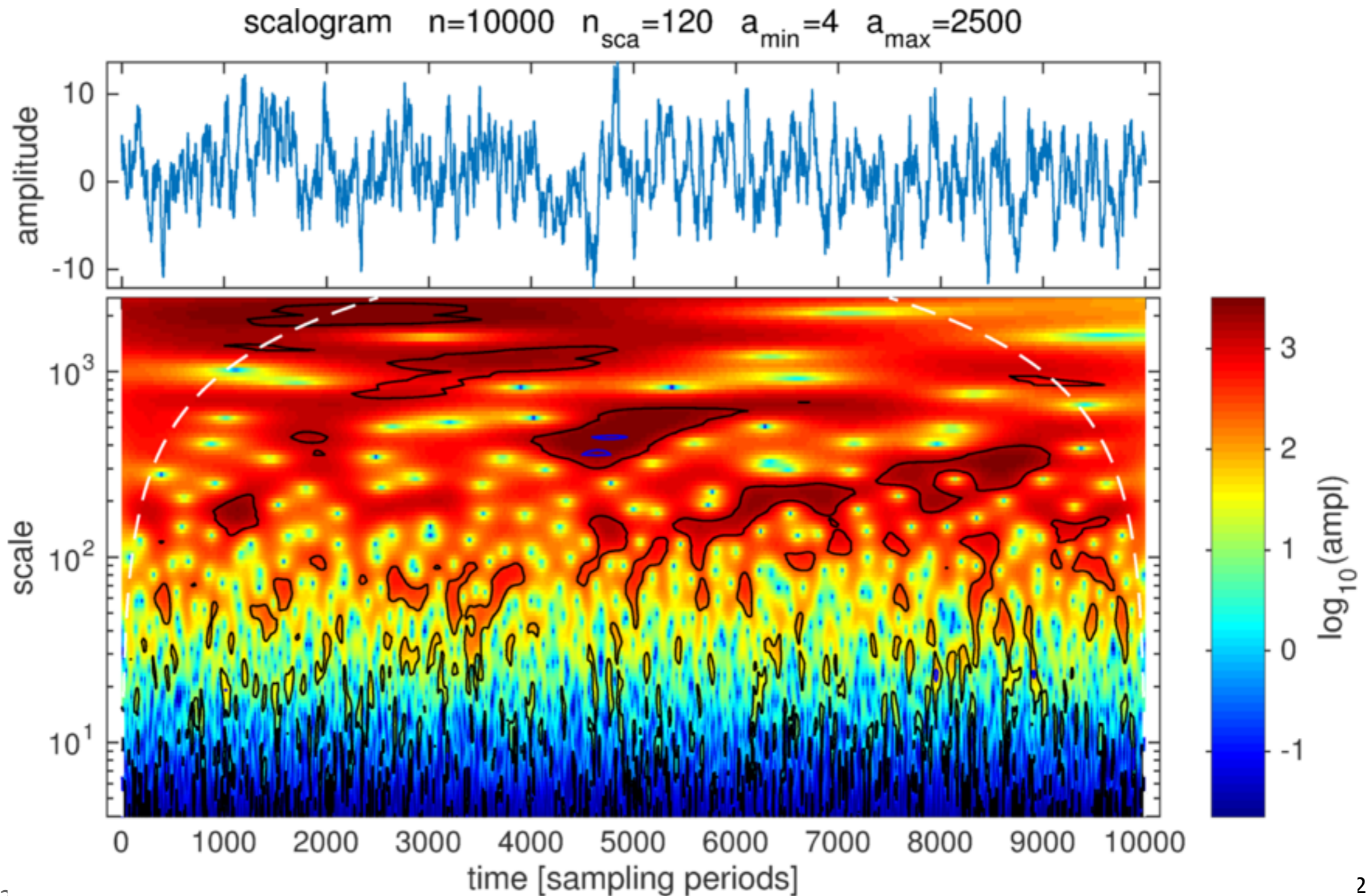
Magnetic field observations

■ Raw data (essentially B_y)



Scalogram

Modulus of continuous wavelet transform (Morlet wavelets)



Intermittency ?

- To identify intermittent bursts of activity : estimate the **local intermittency measure** [Farge, 1992; 1996]

$$B(t) \longrightarrow \tilde{B}(a, \tau) \longrightarrow \text{LIM}(a, \tau) = \frac{|\tilde{B}(a, \tau)|^2}{\langle |\tilde{B}(a, \tau')|^2 \rangle_{\tau'}}$$

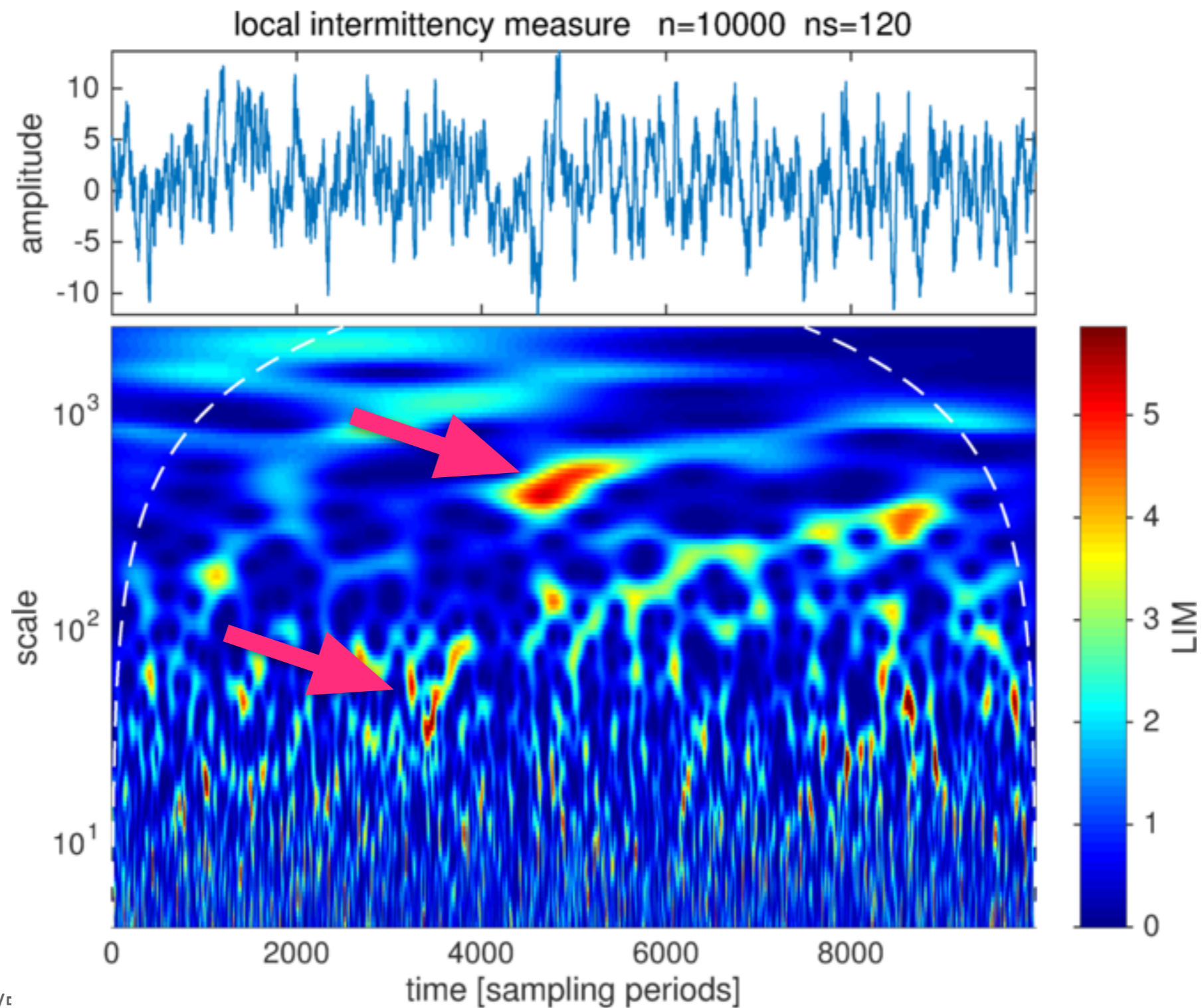
*wavelet transform
at scale a and
time τ*

*local
intermittency
measure*

- Large LIM = local concentration of magnetic energy at scale a and time τ

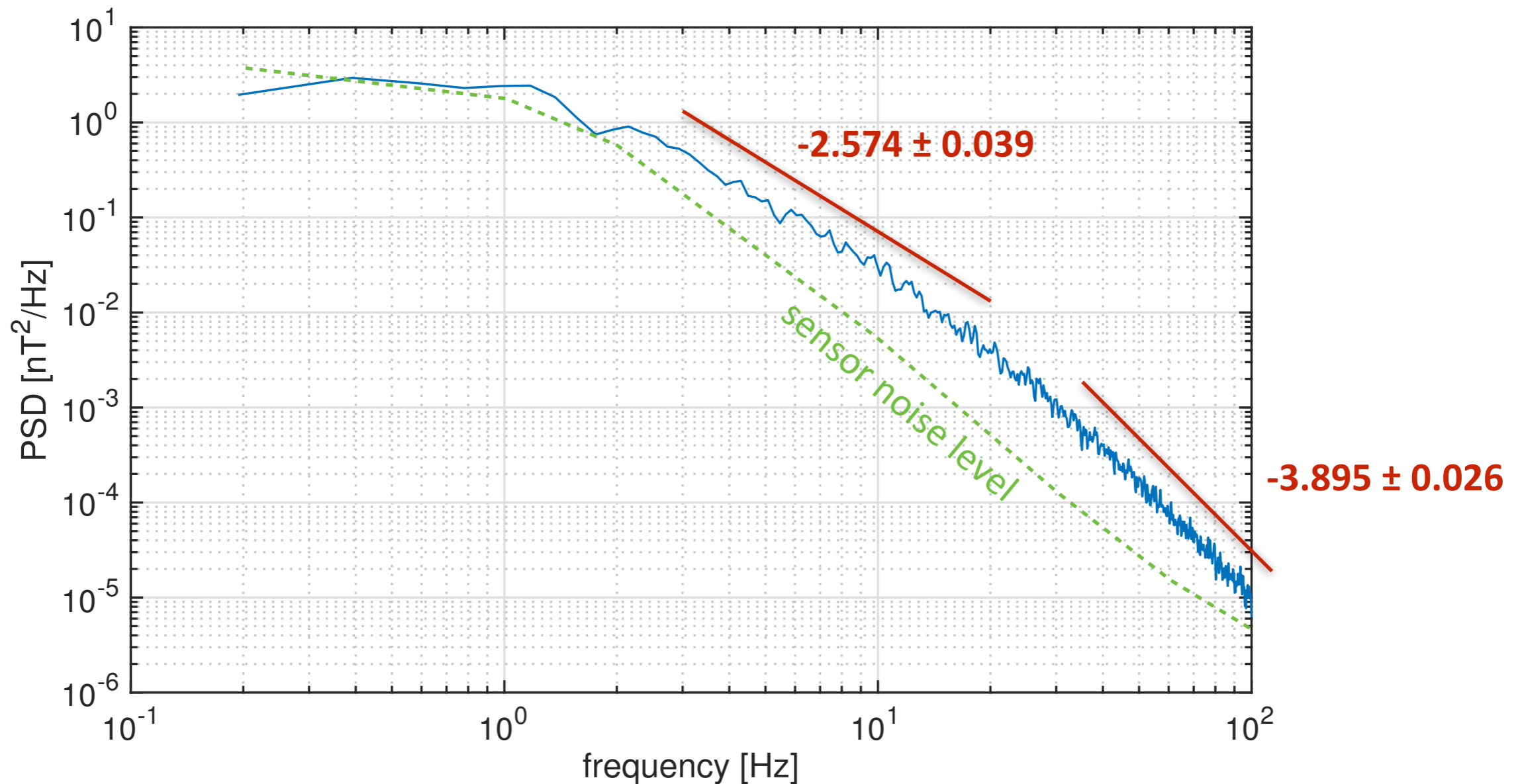
Scalogram

Local intermittency measure



Power spectral density

- Clear evidence for a double power law, with a break



dream

All of this is just a dream

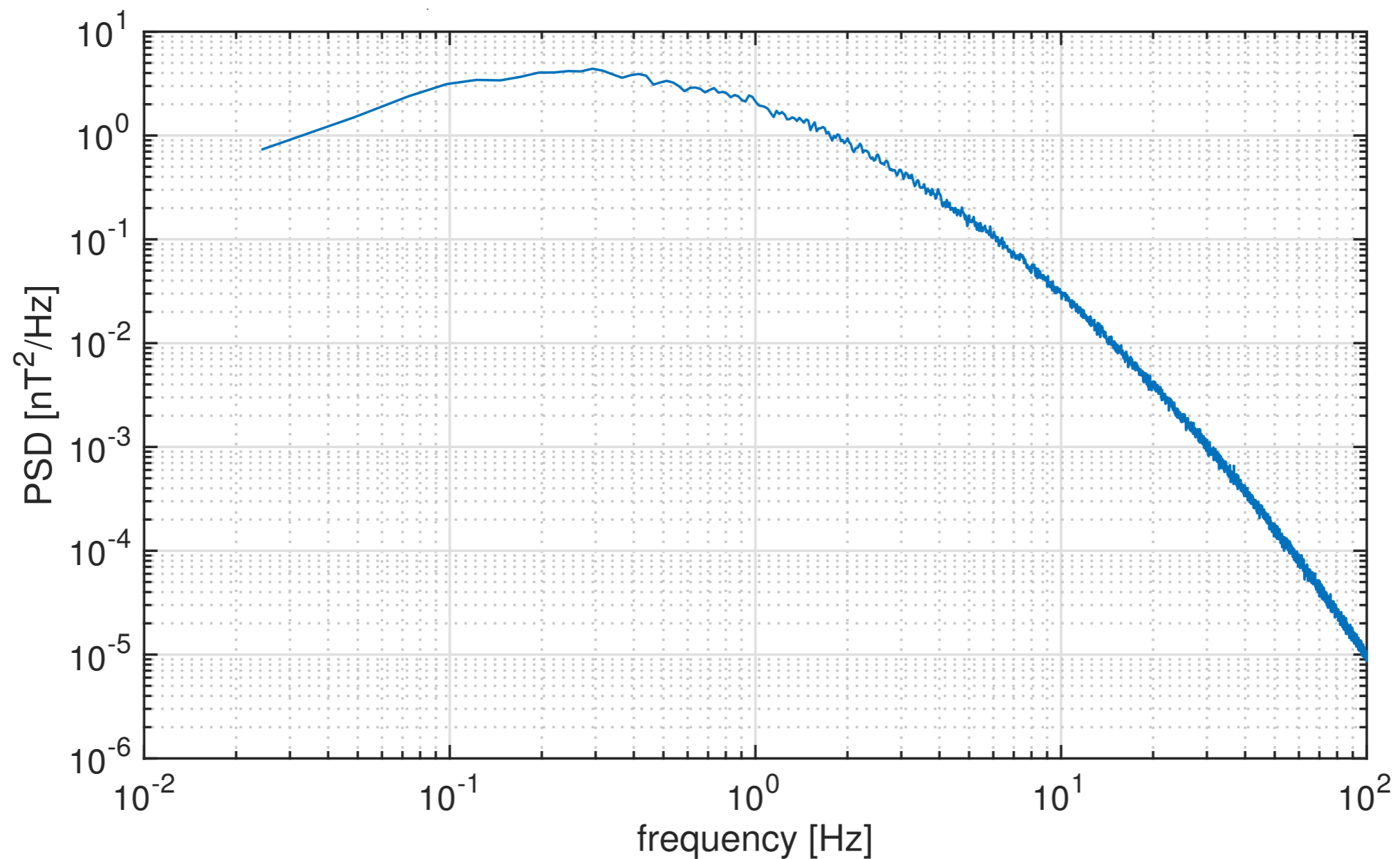


Disclaimer

- My record was generated by a simple stochastic (autoregressive) process:
no intermittency or self-similarity whatsoever!
- The true power spectral density is a log-normal one:
no power laws here!

Disclaimer

- The true power spectral density $\tilde{y}(f) = P_0 \exp\left(-\frac{(f - f_0)^2}{\Delta f^2}\right)$





Intermediate conclusion

Intermediate conclusion

We are completely fooled by our frenzy to detect straight lines

The problem

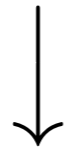
relatively easy

1) Estimate the power spectral density $|\tilde{y}(f)|^2 \propto f^{-\alpha}$



**need a good estimator
of the slope α**

2) Estimate the slope (spectral index) α



**test against a null hypothesis
(« no power law model »)**

3) Prove that the power law model is indeed appropriate

difficult

2 cases occur in practice

case 1 : **data = series of events:**

their **probability density function** follows a power law
(e.g. solar flare distribution)

case 2 : **data = time series:**

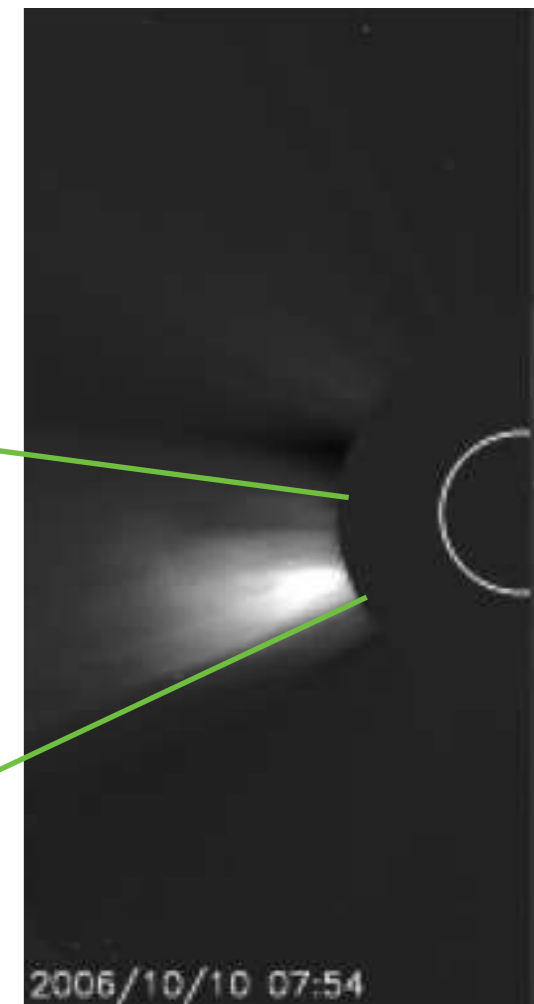
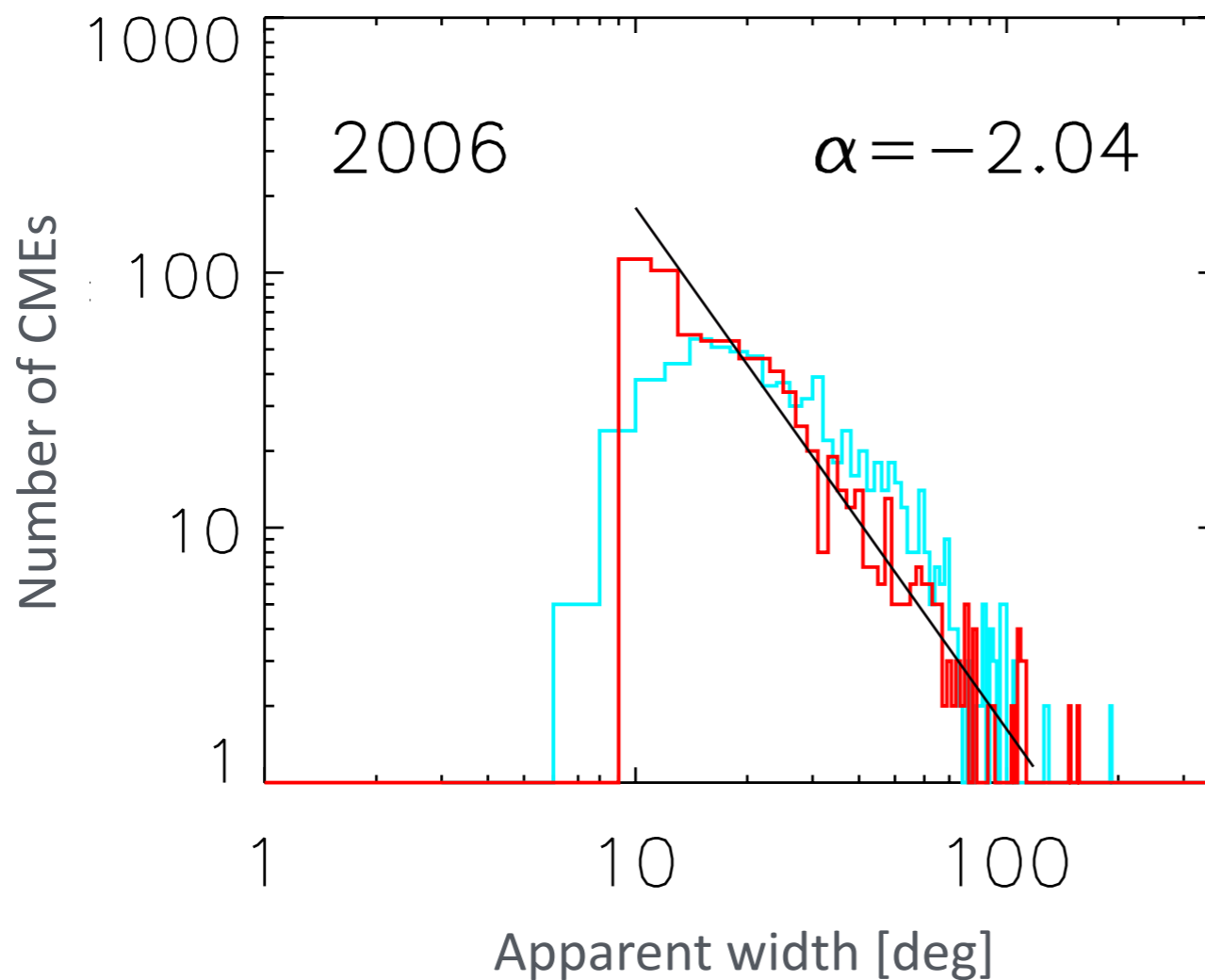
their **power spectral density** follows a power law
(e.g. E or B field record)

The background of the image is a dense, textured field of green, rounded shapes that resemble small plants or biological cells. The shapes are arranged in a somewhat regular, grid-like pattern, with each shape having a slightly raised, bumpy surface. The overall color is a vibrant green, and the lighting is even across the field.

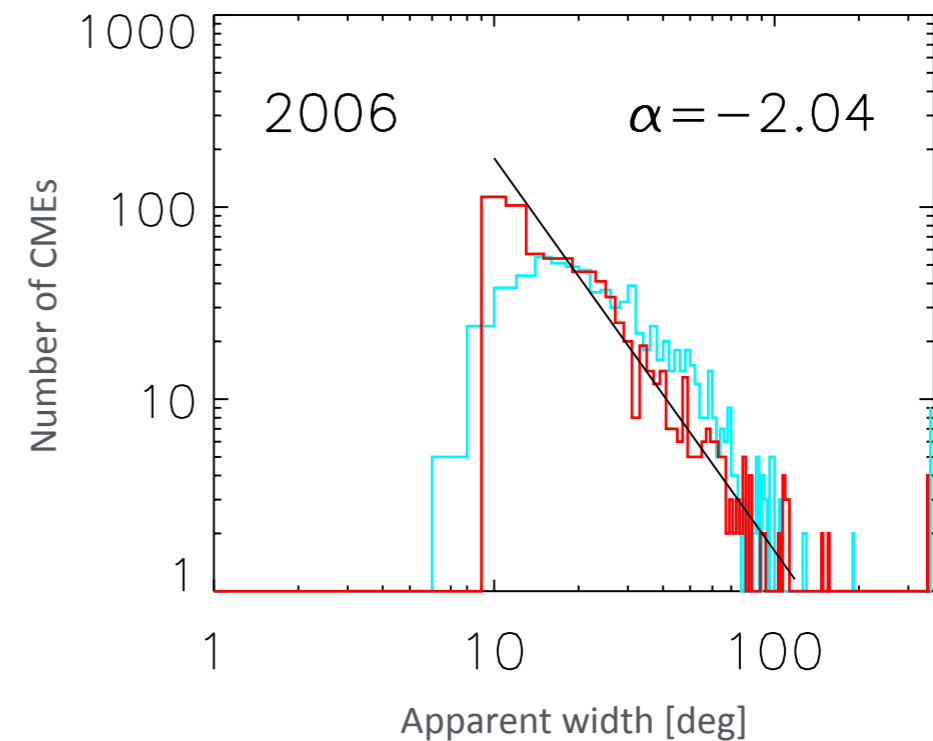
Case 1

Case 1 : the data consist of events

- Example : number of Coronal Mass Ejections (CMEs) per apparent angular width [Robbrecht et al., ApJ 2009]



Case 1 : the data consist of events



■ Classical approach

- bin events to get a histogram (what bin size ?)
- fit by least-squares a line to $\log(\# \text{ of events})$ vs $\log(\text{width})$

This approach is full of flaws and will give biased and noisy estimates. Avoid it!

Case 1 : the data consist of events

- A much better estimator (unbiased, and low variance)

[Clauset et al., SIAM review 2006]

- select the range over which a power law will be fitted

$$y_{min} \leq y_i \leq y_{max}$$

- use the maximum likelihood estimator of the slope

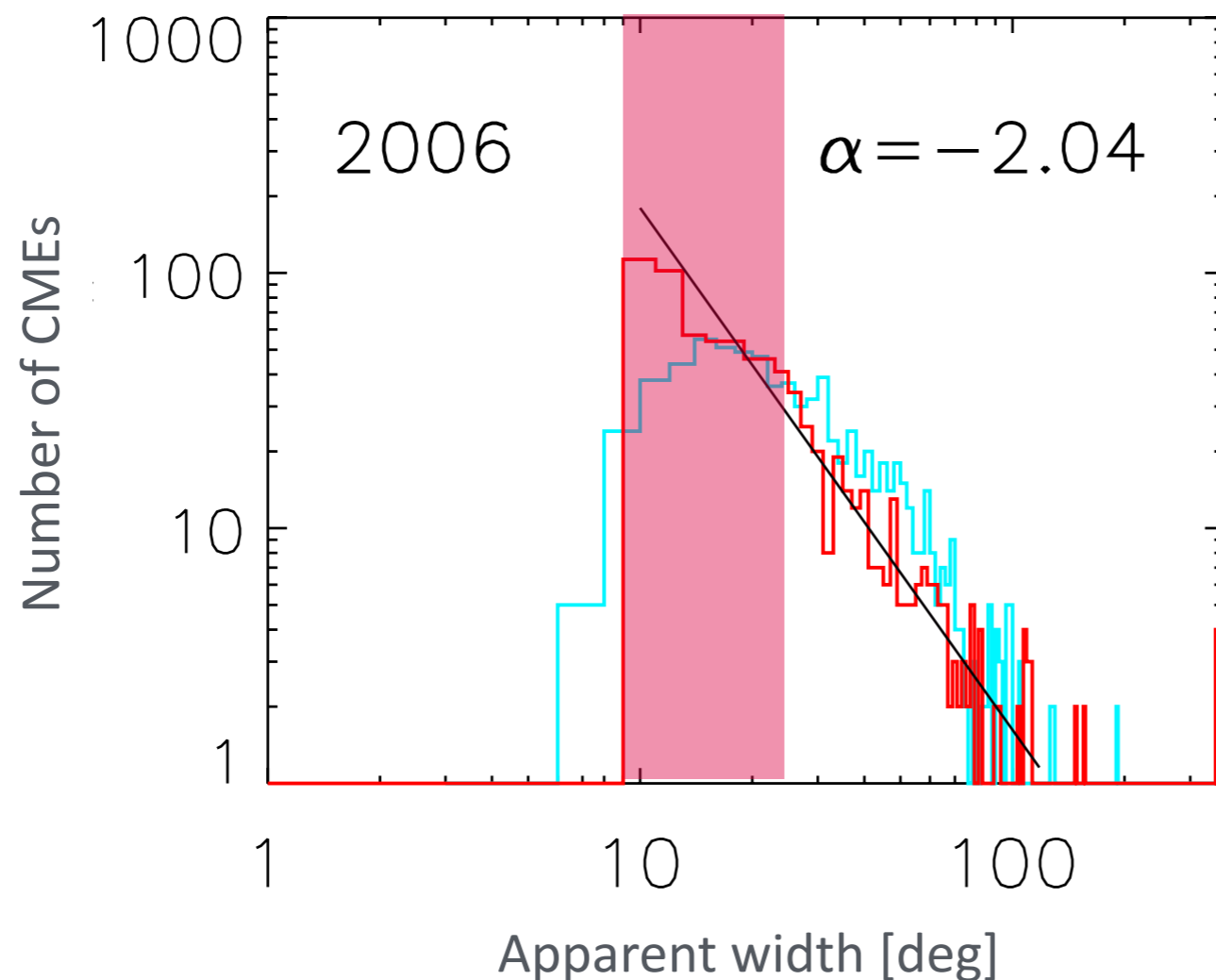
$$\hat{\alpha} = 1 + n \left(\sum_{i=1}^n \log \frac{y_i}{y_{min}} \right)^{-1}$$

- whose standard error is

$$\sigma_{\hat{\alpha}} = \frac{\hat{\alpha} - 1}{\sqrt{n}} + \mathcal{O}(1/n)$$

Case 1 : the data consist of events

- Where does the lower bound (y_{\min}) occur ?



Underestimating the lower bound is far worse than overestimating it

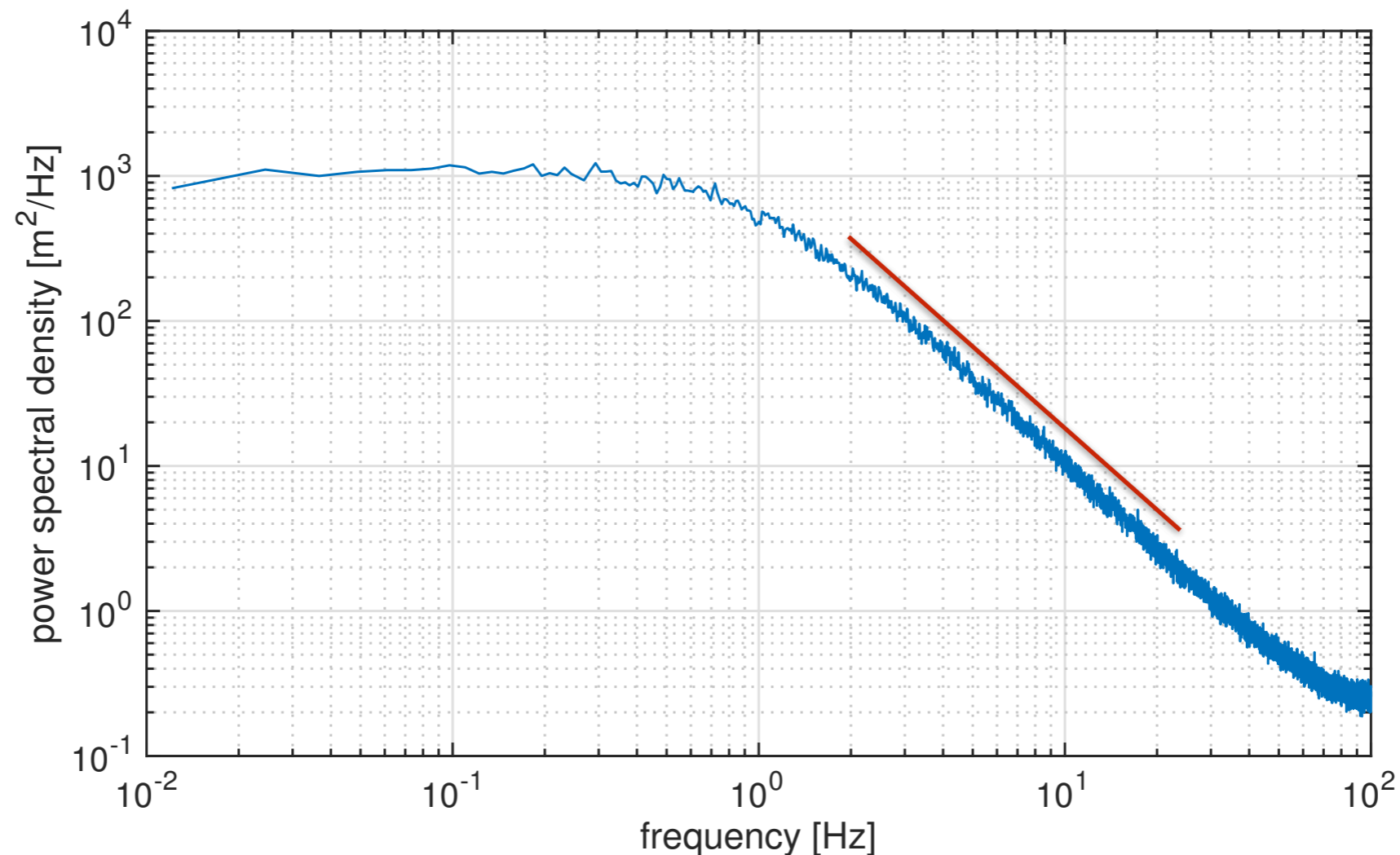
Strategy: select various lower bounds, fit power laws, and determine which one matches the data best

The background of the image is a dense, textured field of green, rounded shapes that resemble small plants or biological cells. The shapes are arranged in a somewhat regular, grid-like pattern, with each shape having a slightly raised, bumpy surface. The overall color is a vibrant green, and the lighting is even across the field.

Case 2

Case 2 : time series

- 2 important steps for time series
 - 1) Estimate the spectral content of the time series
 - 2) Fit power law to the power spectral density



Case 2: time series

- Classical Fourier analysis **is not optimal**:
a time series with self-similarity should be projected on self-similar basis functions (\neq sines)
- Discrete wavelet transform is more appropriate:
self-similar by construction

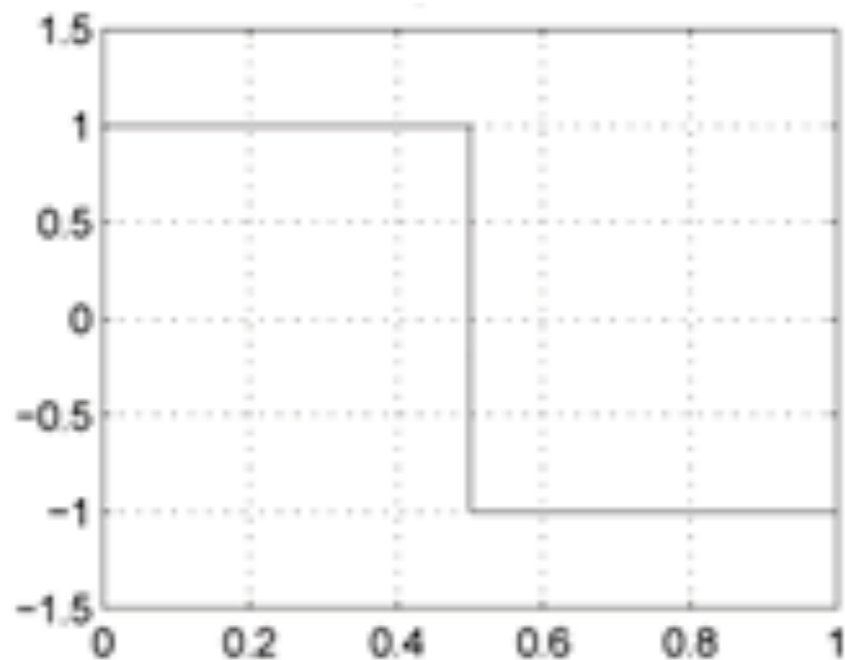
$$\tilde{y}_{j,k} = \int_{\mathbb{R}} y(t) \psi_{j,k}(t) dt, \quad j, k \in \mathbb{Z}$$

$$\text{where } \psi_{j,k}(t) = \frac{1}{2^{j/2}} \psi \left(\frac{t}{2^j} - k \right), \quad j, k \in \mathbb{Z}$$

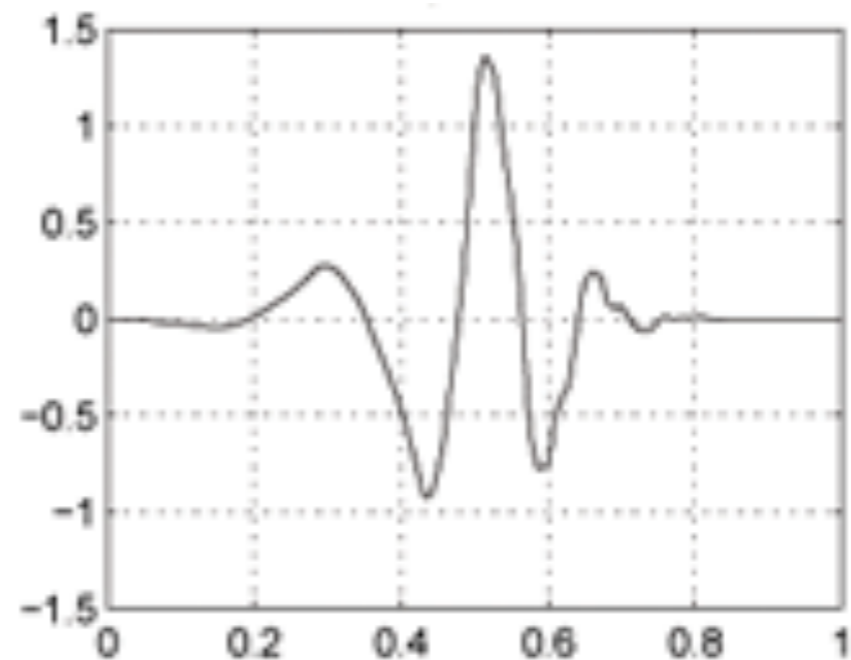
$\psi(\)$ is the mother wavelet

Case 2: time series

- Lots of different mother wavelets around



Haar wavelet (db1)
1 vanishing moment



Daubechies db4 wavelet
4 vanishing moments

- The number of vanishing moments is a crucial parameter

Case 2: time series

Fourier

- Compute transform

$$\tilde{y}(f) = \int y(t) e^{-j2\pi ft} dt$$

- Power spectral density

$$S(f) = |\tilde{y}(f)|^2$$

- can be estimated at all frequencies
- nearby values of $S(f)$ are not independent

Discrete wavelet transform

- Compute transform

$$\tilde{y}_{j,k} = \int_{\mathbb{R}} y(t) \psi_{j,k}(t) dt$$

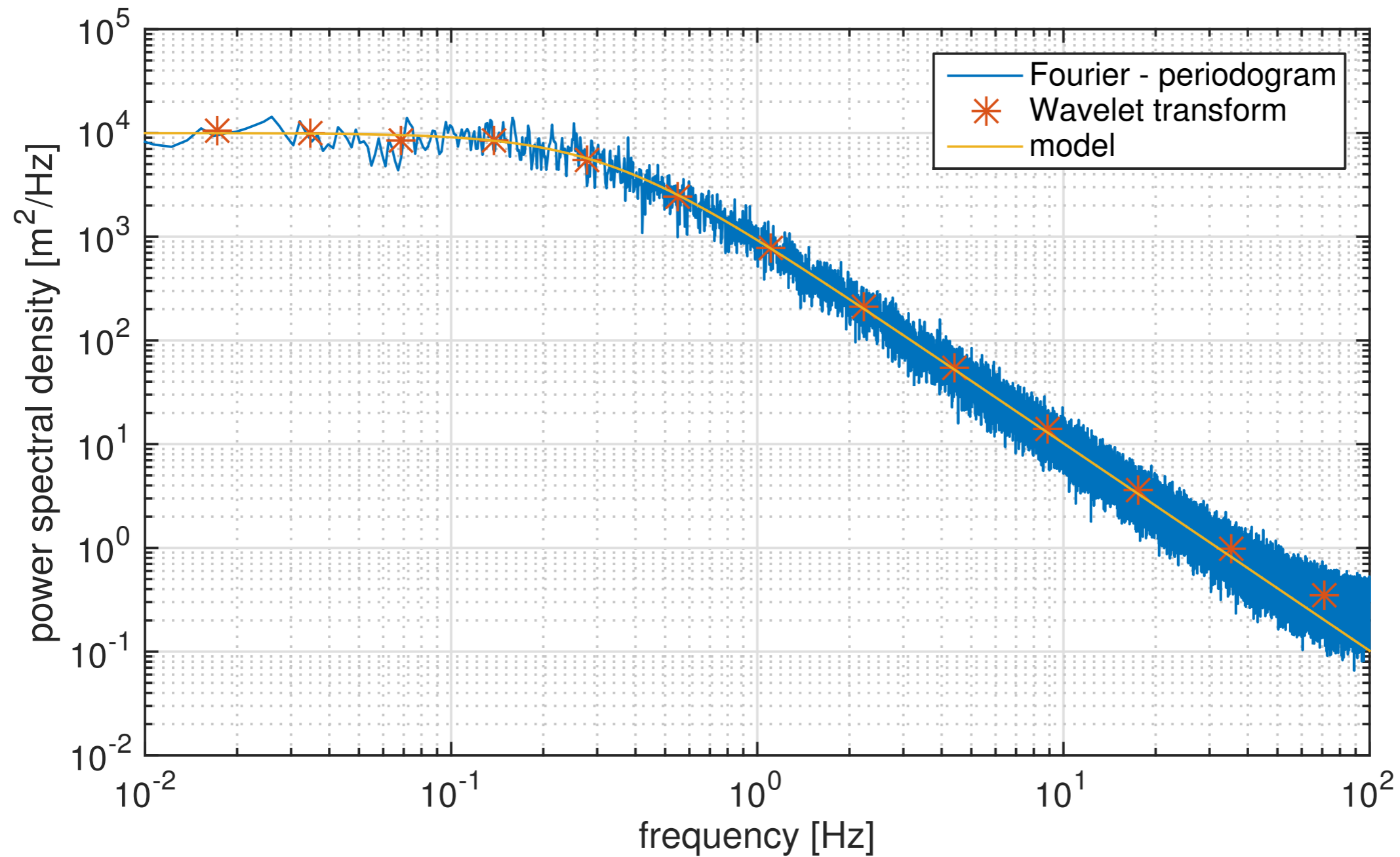
- Power spectral density

$$S_j = \langle |\tilde{y}_{j,k}|^2 \rangle_k$$

- can be estimated per octave only (scale_j = 1, 2, 4, 8, ...)
- wavelet coefficients at nearby scales/positions are independent (very useful)

Case 2: time series

■ Example



Why the order of the wavelet is important

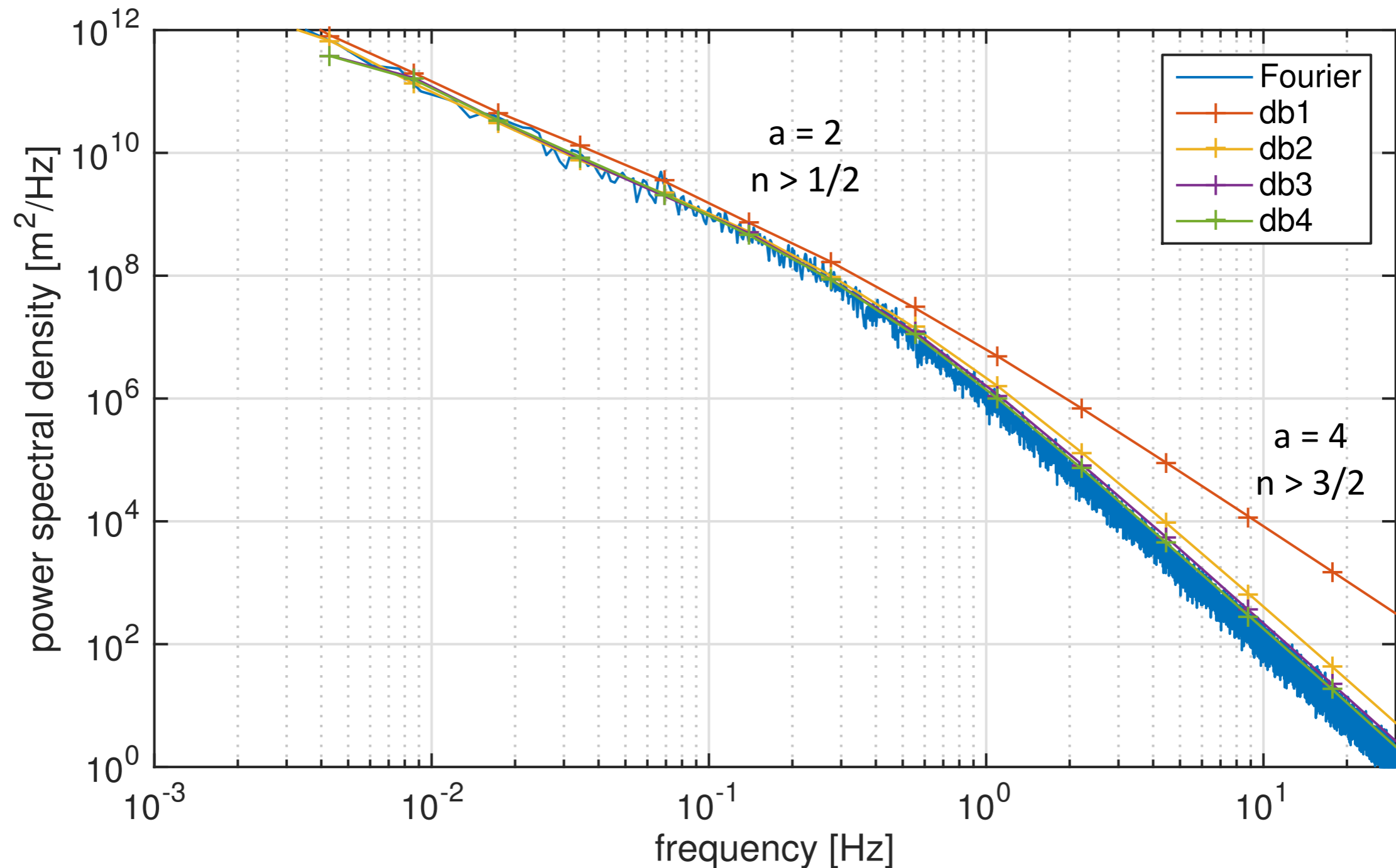
If the mother wavelet does not have enough vanishing moments n

$$\int t^n \psi(t) dt = 0$$



- The wavelet coefficients ($|\tilde{y}_{j,k}|^2$) are not independent anymore
- The estimate of the spectral index will be biased

Why the order of the wavelet is important



For estimating a power law $S(f) \sim f^{-a}$, the number of vanishing moments must be $n > (a-1)/2$



Conclusion

(not a fake one)

Conclusions (1/2)



1. almost any gentle curve will look like a line on a log-log plot
2. plotting scalograms, local intermittency measures, power laws, etc. without proper statistical testing is nonsense.

Conclusions (2/2)

■ For sampled data

- Use maximum likelihood estimator
- Proper estimation of upper and lower (y_{\min}) bounds is crucial

■ For time series

- Use discrete wavelet transform
- Ensure that the order of the wavelet is high enough (stay away from Haar)
- Power law should cover at least a decade to make sense.

■ Not addressed here:

- Use statistical tests (e.g. Kolmogorov-Smirnov) to compare power law model against other plausible models.
- Or better, use a bayesian approach (computationally expensive)

Further reading

- P. Abry, P. Flandrin, M. S. Taqqu, and D. Veitch, *Self-similarity and long-range dependence through the wavelet lens*, Birkhäuser, 2003, 527–556.
- A. Clauset et al., *Power-law distributions in empirical data*, SIAM Review 51 (2009), 661-773
- Y. Virkar and A. Clauset, *Power-law distributions in binned empirical data*, Annals of Applied Statistics, 8 (2014), 89–119.